

# **Computational Intelligence**

Winter Term 2016/17

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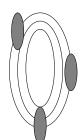
Fakultät für Informatik

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# **Towards CMA-ES**

Lecture 11

claim: mutations should be aligned to isolines of problem (Schwefel 1981)



if true then covariance matrix should be inverse of Hessian matrix!

 $\Rightarrow$  assume f(x)  $\approx \frac{1}{2}$  x'Ax + b'x + c  $\Rightarrow$  H = A

$$Z \sim N(0, C)$$
 with density 
$$f_Z(x) = \frac{1}{(2\pi)^{n/2} |C|^{1/2}} \exp\left(-\frac{1}{2}x'C^{-1}x\right)$$

since then many proposals how to adapt the covariance matrix

 $\Rightarrow$  extreme case: use n+1 pairs (x, f(x)), apply multiple linear regression to obtain estimators for A, b, c invert estimated matrix A! OK, **but**: O(n<sup>6</sup>)! (Rudolph 1992)

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### Lecture 11

mutation: Y = X + Z

Z ~ N(0, C) multinormal distribution

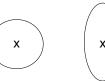
maximum entropy distribution for support R<sup>n</sup>, given expectation vector and covariance matrix

how should we choose covariance matrix C?

unless we have not learned something about the problem during search

⇒ don't prefer any direction!

 $\Rightarrow$  covariance matrix C = I<sub>n</sub> (unit matrix)





 $C = diag(s_1,...,s_n)$ 

C orthogonal

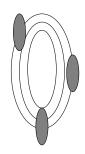
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# **Towards CMA-ES**

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Lecture 11

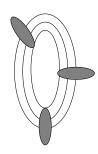
doubts: are equi-aligned isolines really optimal?



principal axis

should point into negative gradient direction!

(proof next slide)



most (effective) algorithms behave like this:

run roughly into negative gradient direction, sooner or later we approach longest main principal axis of Hessian,

now negative gradient direction coincidences with direction to optimum, which is parallel to longest main principal axis of Hessian, which is parallel to the longest main principal axis of the inverse covariance matrix

(Schwefel OK in this situation)

### **Towards CMA-ES**

# Lecture 11

 $Z = rQu, A = B'B, B = Q^{-1}$ 

$$f(x + rQu) = \frac{1}{2}(x + rQu)'A(x + rQu) + b'(x + rQu) + c$$

$$= \frac{1}{2}(x'Ax + 2rx'AQu + r^2u'Q'AQu) + b'x + rb'Qu + c$$

$$= f(x) + rx'AQu + rb'Qu + \frac{1}{2}r^2u'Q'AQu$$

$$= f(x) + r(Ax + b + \frac{r}{2}AQu)'Qu$$

$$= f(x) + r(\nabla f(x) + \frac{r}{2}AQu)'Qu$$

$$= f(x) + r\nabla f(x)'Qu + \frac{r^2}{2}u'Q'AQu$$

$$= f(x) + r\nabla f(x)'Qu + \frac{r^2}{2}u'Q'AQu$$

if Qu were deterministic ...

 $\Rightarrow$  set Qu =  $-\nabla f(x)$  (direction of steepest descent)



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# **Towards CMA-ES**

# Lecture 11

#### Theorem

A quadratic matrix  $C^{(k)}$  is symmetric and positive definite for all  $k \ge 0$ , if it is built via the iterative formula  $C^{(k+1)} = \alpha_k C^{(k)} + \beta_k v_k v_k'$  where  $C^{(0)} = I_n$ ,  $v_k \ne 0$ ,  $\alpha_k > 0$  and liminf  $\beta_k > 0$ .

#### **Proof:**

If  $v \neq 0$ , then matrix V = vv' is symmetric and positive semidefinite, since

- as per definition of the dyadic product  $v_{ij} = v_i \cdot v_j = v_i \cdot v_j = v_{ij}$  for all i, j and
- for all  $x \in \mathbb{R}^n$ :  $x'(vv') x = (x'v) \cdot (v'x) = (x'v)^2 \ge 0$ .

Thus, the sequence of matrices  $v_k v_k'$  is symmetric and p.s.d. for  $k \ge 0$ .

Owing to the previous lemma matrix C(k+1) is symmetric and p.d., if

 $C^{(k)}$  is symmetric as well as p.d. and matrix  $v_k v_k'$  is symmetric and p.s.d.

Since  $C^{(0)} = I_n$  symmetric and p.d. it follows that  $C^{(1)}$  is symmetric and p.d.

Repetition of these arguments leads to the statement of the theorem.

#### **Towards CMA-ES**

Lecture 11

# Apart from (inefficient) regression, how can we get matrix elements of Q?

 $\Rightarrow$  iteratively:  $C^{(k+1)} = \text{update}(C^{(k)}, Population}^{(k)})$ 

<u>basic constraint</u>:  $C^{(k)}$  must be positive definite (p.d.) and symmetric for all  $k \ge 0$ ,

otherwise Cholesky decomposition impossible: C = Q'Q

#### Lemma

Let A and B be quadratic matrices and  $\alpha$ ,  $\beta > 0$ .

- a) A, B symmetric  $\Rightarrow \alpha A + \beta B$  symmetric.
- b) A positive definite and B positive semidefinite  $\Rightarrow \alpha$  A +  $\beta$  B positive definite

#### Proof:

ad a) C = 
$$\alpha$$
 A +  $\beta$  B symmetric, since  $c_{ij}$  =  $\alpha$   $a_{ij}$  +  $\beta$   $b_{ij}$  =  $\alpha$   $a_{ji}$  +  $\beta$   $b_{ji}$  =  $c_{ji}$ 

ad b) 
$$\forall x \in \mathbb{R}^n \setminus \{0\}$$
:  $x'(\alpha A + \beta B) x = \underbrace{\alpha x'Ax}_{>0} + \underbrace{\beta x'Bx}_{\geq 0} > 0$ 

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Lecture 11

# CMA-ES

**Idea:** Don't estimate matrix **C** in each iteration! Instead, approximate <u>iteratively!</u>

→ Covariance Matrix Adaptation Evolutionary Algorithm (CMA-EA)

Set initial covariance matrix to  $C^{(0)} = I_n$ 

$$C^{(t+1)} = (1-\eta) C^{(t)} + \eta \sum_{i=1}^{\mu} w_i (x_{i:\lambda} - m^{(t)}) (x_{i:\lambda} - m^{(t)})'$$

 $\eta~$  : "learning rate"  $\in$  (0,1)

(Hansen, Ostermeier et al. 1996ff.)

 $w_{i}$ : weights; mostly  $1/\mu$ 

$$m = \frac{1}{\mu} \sum_{i=1}^{\mu} x_{i:\lambda}$$
 mean of all selected parents

complexity:  $O(\mu n^2 + n^3)$ 

sorting:  $f(x_{1:\lambda}) \le f(x_{2:\lambda}) \le ... \le f(x_{\lambda:\lambda})$ 

Caution: must use mean m(t) of "old" selected parents; not "new" mean m(t+1)!

⇒ Seeking covariance matrix of fictitious distribution pointing in gradient direction!

# **CMA-ES** Lecture 11

State-of-the-art: CMA-EA (currently many variants)

→ many successful applications in practice

# available in WWW:

Fortran, Python, Matlab, R, Scilab • <a href="http://www.lri.fr/~hansen/cmaes\_inmatlab.html">http://www.lri.fr/~hansen/cmaes\_inmatlab.html</a>

 http://shark-project.sourceforge.net/ (EAlib, C++)

• ...

# advice:

before designing your own new method or grabbing another method with some fancy name ... try CMA-ES – it is available in most software libraries and often does the job!



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C, C++, Java

